Module 5: Introduction to Multilevel Modelling MLwiN Practicals

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Pre-requisites

• Modules 1-4

Contents

P5.1	Com	nparing Groups using Multilevel Modelling	. 3
P5.1 P5.1	.1 .2	A multilevel model of attainment with school effects Examining school effects (residuals)	.3 .5
P5.2	Add	ing Student-level Explanatory Variables: Random Intercept Models	. 8
P5.3	Allo	wing for Different Slopes across Schools: Random Slope Models	11
P5.3 P5.3 P5.3 P5.3 P5.3 P5.3	8.1 8.2 8.3 8.4 8.5 8.6	Testing for random slopes Interpretation of random cohort effects across schools Examining intercept and slope residuals for schools Between-school variance as a function of cohort Adding a random coefficient for gender (dichotomous x) Adding a random coefficient for social class (categorical x)	12 12 13 15 17 19
P5.4	Add	ing Level 2 Explanatory Variables	24
P5.4 P5.4	1.1 1.2	Contextual effects Cross-level interactions	25 28
P5.5	Com	nplex Level 1 Variation	31
P5.5 P5.5 P5.5	5.1 5.2 5.3	Within-school variance as a function of cohort (continuous X) Within-school variance as a function of gender (dichotomous X) Within-school variance as a function of cohort and gender	32 34 37

Some of the sections within this module have online quizzes for you to test your understanding. To find the quizzes:

EXAMPLE

From within the LEMMA learning environment

- Go down to the section for Module 5: Introduction to Multilevel Modelling
- Click "5.1 Comparing Groups Using Multilevel Modelling" to open Lesson 5.1
- Click Q1 to open the first question

Introduction to the Scottish Youth Cohort Trends Dataset

You will be analysing data from the Scottish School Leavers Survey (SSLS), a nationally representative survey of young people. We use data from seven cohorts of young people collected in the first sweep of the study, carried out at the end of the final year of compulsory schooling (aged 16-17) when most sample members had taken Standard grades1.

In the practical for Module 3 on multiple regression, we considered the predictors of attainment in Standard grades (subject-based examinations, typically taken in up to eight subjects). In this practical, we extend the (previously single-level) multiple regression analysis to allow for dependency of exam scores within schools and to examine the extent of between-school variation in attainment. We also consider the effects on attainment of several school-level predictors.

The dependent variable is a total attainment score. Each subject is graded on a scale from 1 (highest) to 7 (lowest) and, after recoding so that a high numeric value denotes a high grade, the total is taken across subjects.

The analysis dataset contains the student-level variables considered in Module 3 together with a school identifier and three school-level variables:

Variable name	Description and codes
CASEID	Anonymised student identifier
SCHOOLID	Anonymised school identifier

¹ We are grateful to Linda Croxford (Centre for Educational Sociology, University of Edinburgh) for providing us with these data. The dataset was constructed as part of an ESRC-funded project on Education and Youth Transitions in England, Wales and Scotland 1984-2002.

Further analyses of the data can be found in Croxford, L. and Raffe, D. (2006) "Education Markets and Social Class Inequality: A Comparison of Trends in England, Scotland and Wales". In R. Teese (Ed.) *Inequality Revisited*. Berlin: Springer.

SCORE	Point score calculated from awards in Standard grades taken a age 16. Scores range from 0 to 75, with a higher score indicating a higher attainment										
COHORT90	The sample includes the following cohorts: 1984, 1986, 1988, 1990, 1996 and 1998. The COHORT90 variable is calculated by subtracting 1990 from each value. Thus values range from -6 (corresponding to 1984) to 8 (1998), with 1990 coded as zero										
FEMALE	Sex of student (1=female, 0=male)										
SCLASS	Social class, defined as the higher class of mother or father (1=managerial and professional, 2=intermediate, 3=working, 4=unclassified).										
SCHTYPE	School type, distinguishing independent schools from state- funded schools (1=independent, 0=state-funded)										
SCHURBAN	Urban-rural classification of school (1=urban, 0=town or rural)										
SCHDENOM	School denomination (1=Roman Catholic, 0=non- denominational)										

There are 33988 students in 508 schools.

Open the worksheet to

From within the LEMMA Learning Environment

- Go to Module 5: Introduction to MultilevelModelling, and scroll down to MLwiN Datafiles
- If you do not already have MLwiN to open the datafile with, click(get MLwiN).
- Click "
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 ⁵
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You will see the **Names** window:

	Names										- • ×							
Column:	Name	Description	Toggle Categorica	Data:	View	Сору	Paste	Delete	Cate	gories: \	View	Сору	Paste	Regenerate	Window:	Used columns	0	Help
Name		Cn	n	missing	min		max	catego	orical	descriptio	on							^
caseid		1	33988	0	1		38192	False										
schoolid		2	33988	0	1		511	False										
score		3	33988	0	0		75	False										
cohort90		4	33988	0	-6		8	False										
female		5	33988	0	0		1	False										
sclass		6	33988	0	1		4	False										
schtype		7	33988	0	0		1	False										
schurban		8	33988	0	0		1	False										
schdenom	1	9	33988	0	0		1	False										
C10		10	0	0	0		0	False										
C11		11	0	0	0		0	False										
C12		12	0	0	0		0	False										
C13		13	0	0	0		0	False										
C14		14	0	0	0		0	False										
C15		15	0	0	0		0	False										

P5.1 Comparing Groups using Multilevel Modelling

P5.1.1 A multilevel model of attainment with school effects

We will start with the simplest multilevel model which allows for school effects on attainment, but without explanatory variables. This 'null' model may be written

$$y_{ij} = \beta_0 + u_j + e_{ij}$$
 (5.1)

where y_{ij} is the attainment of student *i* in school *j*, β_0 is the overall mean across schools, u_j is the effect of school *j* on attainment, and e_{ij} is a student-level residual. The school effects u_j , which we will also refer to as school (or level 2) residuals, are assumed to follow a normal distribution with mean zero and variance σ_u^2 .

To set up this model in MLwiN:

- From the Model menu, select Equations
- Click Notation at the bottom of the Equations window, clear the general tick box, and click Done
- Click on y and select SCORE from the drop-down list
- Click on N Levels and select 2-ij
- For level 2(j), select SCHOOLID
- For level 1(i), select CASEID
- Click Done
- Click on β_0 , and check **j(schoolid)** to introduce a random school effect, and click **Done**. Click + and notice that this step leads to the addition of u_{0j} to the model
- Click +again to see the full model specification

The model should look like this:

Equations	- - ×								
$\text{score}_{ij} = \beta_{0j} + e_{ij}$									
$\beta_{0j} = \beta_0 + u_{0j}$									
$u_{0j} \sim N(0, \sigma_{u0}^2)$									
$\boldsymbol{e}_{ij} \sim \mathrm{N}(0, \sigma_{\boldsymbol{e}}^2)$									
(33988 of 33988 cases in use)									
UNITS:									
schoolid: 508 (of 508) in use									
Name + - Add Term Estimates Nonlinear Clear Notation Responses Store Help Zoom 100	0 -								

If the second equation $(\beta_{0j} = \beta_0 + u_{0j})$ is substituted for β_{0j} in the first equation, we obtain a equation that has the same form as (5.1).

Notice that a '0' subscript has been added to the school effect u_j and its variance σ_u^2 , in anticipation of adding further random effects at the school level (see P5.3).

- Click **Start** to fit the model
- Click Estimates twice to see the parameter estimates

The overall mean attainment (across schools) is estimated as 30.60. The mean for school *j* is estimated as 30.60 + \hat{u}_{0j} , where \hat{u}_{0j} is the school residual which we will estimate in a moment. A school with \hat{u}_{0j} >0 has a mean that is higher than average, while \hat{u}_{0j} <0 for a below-average school. (We will obtain confidence intervals for residuals to determine whether differences from the overall mean can be considered 'real' or due to chance.)

Partitioning variance

The between-school (level 2) variance in attainment is estimated as $\hat{\sigma}_{u0}^2$ =60.99, and the within-school between-student (level 1) variance is estimated as $\hat{\sigma}_e^2$ =258.36. Thus the total variance is 60.99+258.36=319.35.

The variance partition coefficient (VPC) is 60.99/319.35 = 0.19, which indicates that 19% of the variance in attainment can be attributed to differences between schools. Note, however, that we have not accounted for intake ability (measured by exams taken on entry to secondary school) so the school effects are not value-added. Previous studies have found that between-school variance in *progress*, i.e. after accounting for intake attainment, is close to 10%.

Testing for school effects

To test the significance of school effects, we can carry out a likelihood ratio test comparing the null multilevel model with a null single-level model. To fit the null single-level model, we need to remove the random school effect:

- In the Equations window, click on β_{0i}
- Click on the check box next to j(schoolid) to uncheck it, and click Done
- The *u*_{0*i*} should be removed from the model
- Click Start to fit the model

Equationsscore $ij = 31.095(0.094) + e_{ij}$ $e_{ij} \sim N(0, \sigma_e^2) - \sigma_e^2 = 299.779(2.300)$ -2*loglikelihood = 290288.841(33988 of 33988 cases in use)Name + - Add Term Estimates Nonlinear Clear Notation Responses Store Help Zoom 100 •

The likelihood ratio test statistic is calculated as the difference in the -2*loglikelihood values for the two models:

LR = 290289 - 286539 = 3750 on 1 d.f. (because there is only parameter difference between the models, σ_{u0}^2).

Bearing in mind that the 5% point of a chi-squared distribution on 1 d.f. is 3.84, there is overwhelming evidence of school effects on attainment. We will therefore revert to the multilevel model with school effects.

- In the **Equations** window, click on β₀ (or its estimate 31.095)
- Check **j(schoolid)** and click **Done**
- *u*_{0j} should be returned to the model
- Click Start to fit the model

P5.1.2 Examining school effects (residuals)

To estimate the school-level residuals and produce a caterpillar plot:

- From the Model menu, select Residuals
- Select the **Settings** tab of the **Residuals** window
- Next to level: change from 1:caseid to 2:schoolid
- In the text box next to SD(comparative) of residual to edit 1.0 to 1.96, so that we obtain 95% confidence limits
- Click Calc
- Select the Plots tab
- Under single, check residual +/-1.96 sd x rank
- Click Apply

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