

Single-level Models for Binary Responses

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y_i response for individual i ($i = 1, \dots, n$), coded 0 or 1

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$$E(y) = \pi = \Pr(y = 1)$$

$$\text{var}(y) = \pi(1 - \pi)$$

$$\text{estimated by } \hat{\pi} = \frac{r}{n}$$

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y follows a **Bernoulli distribution** (special case of binomial distribution)

Linear Probability Model

Model for the mean of y

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or expressed as a model for y_i

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where $e_i \sim N(0, \sigma_e^2)$

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Estimate using ordinary least squares (as for continuous y).

Problems with the Linear Probability Model

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- Relationship between π and x may be **nonlinear**, although linearity assumption often reasonable for π between 0.2 and 0.8
- Possible to get **predicted probabilities outside $[0,1]$** . Again, this is unlikely if π lies between 0.2 and 0.8 for all x (or combinations of values on a set of x variables)

The Generalised Linear Model

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usually written as

$$F^{-1}(\pi_i) = \beta_0 + \beta_1 x_i$$

where F^{-1} is called the **link function**

The Logit/Logistic Model

Write $z = \beta_0 + \beta_1 x$

Logistic transformation of z

$$\pi = F(z) = \frac{\exp(z)}{1 + \exp(z)} = \frac{e^z}{1 + e^z}$$

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Logit model

$$\log\left(\frac{\pi}{1 - \pi}\right) = z = \beta_0 + \beta_1 x$$

where $\pi/(1 - \pi)$ is the **odds** that $y = 1$ and $\log[\pi/(1 - \pi)]$ is the log-odds or **logit**

Interpretation of Logit Coefficients

Take exponentials of each side of logit model:

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Interpret $\exp(\beta_1)$ as an **odds ratio**, comparing the odds for two individuals with x -values spaced 1 unit apart.

Example: State Differences in US Voting

$y = 1$ if intends to vote Bush in 2004 election, 0 otherwise

Variable	$\hat{\beta}$	$se(\hat{\beta})$	$\exp(\hat{\beta})$
Constant	-0.34	0.05	0.71
State (ref=California)			
New York	-0.19	0.08	0.83
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- Odds of voting Bush in California = 0.71
- Odds of voting Bush in New York are 0.83 times odds in California (17% lower)
- Odds of voting Bush in Texas are twice the odds in California (100% higher)

Other Link Functions

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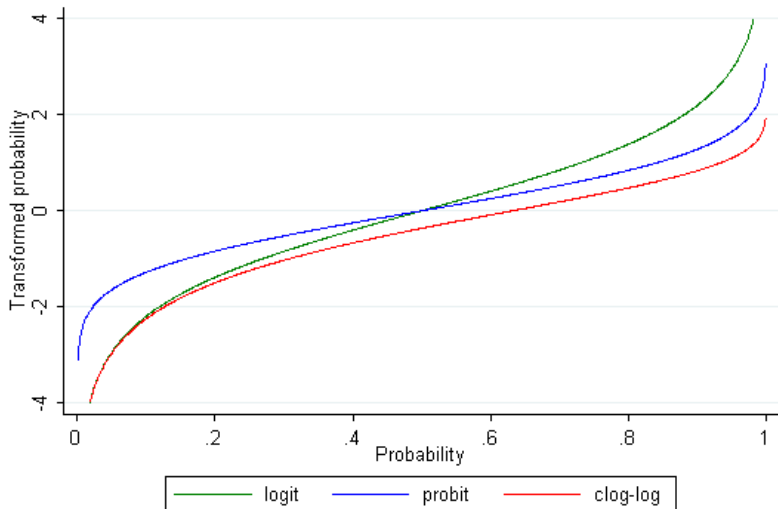
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- In general, get very similar $\hat{\pi}$ whatever link is used
- $\text{Logit}(0.5) = \text{probit}(0.5)$ but move further apart as π gets close to 0 or 1
- Logit and clog-log almost indistinguishable for small π

Logit, probit and clog-log transformations of π



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$$y_i = \begin{cases} 1 & \text{if } y_i^* \geq 0 \\ 0 & \text{if } y_i^* < 0 \end{cases}$$

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- $e_i^* \sim N(0, 1) \rightarrow$ **probit** model
- $e_i^* \sim$ standard logistic (with variance $\simeq 3.29$) \rightarrow **logit** model

Relationship between probit and logit

The residual in the threshold model has fixed variance, but the value it is fixed at depends on the chosen distribution:

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Standard normal	1	probit
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Increasing the variance (and therefore scale of y^*) increases the magnitude of the coefficients.

$$\hat{\beta}_{\text{logit}} \simeq \sqrt{3.29} \hat{\beta}_{\text{probit}} = 1.8 \hat{\beta}_{\text{probit}}$$

US Voting: Logit and Probit Coefficients

Compare coefficients from logit and probit models of voting Bush.

Variable	$\hat{\beta}_{\text{probit}}$	$\hat{\beta}_{\text{logit}}$	$\hat{\beta}_{\text{logit}}/\hat{\beta}_{\text{probit}}$
Constant	-0.21	-0.34	1.62
State (ref=California)			
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But, for this simple model, predicted probabilities of voting Bush will be exactly the same for logit and probit models.

Predicted Probabilities of Voting Bush

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Calculate using NORMDIST in Excel or CDFNORM in SPSS

Significance Testing

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Compare Wald = Z^2 with χ_1^2

Also used to test more general hypotheses, e.g. $H_0 : \beta_1 = \beta_2 = 0$
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Likelihood ratio test

Compare difference in deviance (-2 log-likelihood) between two nested models with χ_q^2 , where q = difference in number of parameters. **Not available in MLwiN for discrete response models.**

Confidence Intervals for β and Odds Ratios

95% CI for β

$$(\hat{\beta} - 1.96 \text{ se}(\hat{\beta}), \hat{\beta} + 1.96 \text{ se}(\hat{\beta}))$$

E.g. 95% CI for NY-California difference in log-odds of voting Bush is $-0.19 \pm (1.96 \times 0.08) = -0.35$ to -0.03

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95% CI for odds ratio, $\exp(\beta)$

Take exponent of lower and upper limits of 95% interval for β .

E.g. 95% CI for NY vs California odds ratio is $\exp(-0.35)$ to $\exp(-0.03) = 0.70$ to 0.97

Adding Further Predictors to Logit Model

Variable	$\hat{\beta}$	$se(\hat{\beta})$	Wald	$\exp(\hat{\beta})$
Constant	-0.42	0.11	-	0.66
State (ref=California)				
New York	-0.18	0.08	4.4	0.84
Texas	0.71	0.08	76.7	2.03
Female	-0.27	0.07	15.9	0.76
Age (years)	0.005	0.002	5.3	1.005

Note: Wald for joint test of state effects is 111.6 on 2 df ($\chi^2_{2;0.05} = 5.99$).
All other tests on 1 df ($\chi^2_{1;0.05} = 3.84$).

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- Odds of voting Bush are 25% lower for women than for men (controlling for state and age)
- Odds of voting Bush increased by 0.5% for each 1-year increase in age (controlling for state and gender)

Predicted Probabilities from Logit Model

Variable	Value	$\hat{\pi}$
State	California	0.42
	New York	0.37
	Texas	0.59
Sex	Male	0.49
	Female	0.42
Age (years)	20	0.42
	30	0.43
	40	0.45

- Values of each variable varied in turn, holding others at sample mean (proportions for categorical variables)
- E.g. $\hat{\pi}$ for State calculated for each possible set of values for the State dummies, with Female=0.54 and Age=46.7 years

Interaction Effects

Suppose we believe the effect of age on voting intentions differs across states, then fit:

$$\begin{aligned} \log\left(\frac{\pi}{1-\pi}\right) &= \beta_0 + \beta_1 \text{ NY} + \beta_2 \text{ TEX} + \beta_3 \text{ FEMALE} + \beta_4 \text{ AGE} \\ &+ \beta_5 \text{ NY} \cdot \text{AGE} + \beta_6 \text{ TEX} \cdot \text{AGE} \end{aligned}$$

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$$\log\left(\frac{\pi}{1-\pi}\right) = \beta_0 + \beta_1 \text{ NY} + \beta_2 \text{ TEX} + \beta_3 \text{ FEMALE} + \beta_4 \text{ AGE} \\ + \beta_5 \text{ NY} \cdot \text{AGE} + \beta_6 \text{ TEX} \cdot \text{AGE}$$

State	Age effect	Estimated age effect
California	β_4	0.005
New York	$\beta_4 + \beta_5$	0.005+0.010
Texas	$\beta_4 + \beta_6$	0.005-0.012

Interaction Effects

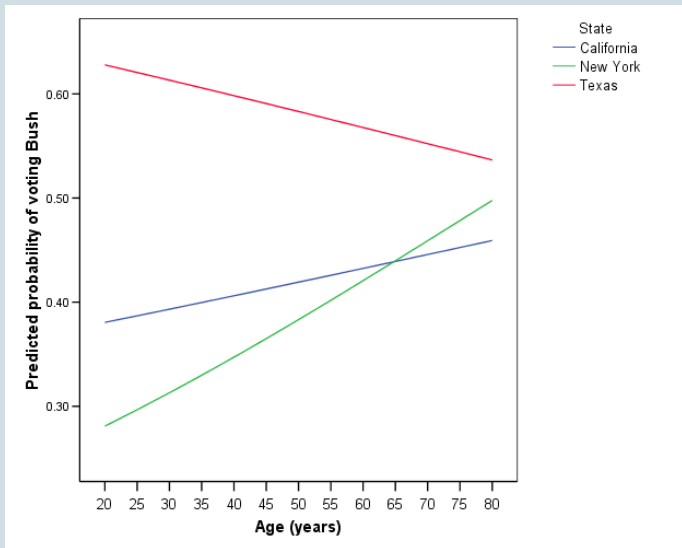
Suppose we believe the effect of age on voting intentions differs across states, then fit:

$$\log\left(\frac{\pi}{1-\pi}\right) = \beta_0 + \beta_1 \text{ NY} + \beta_2 \text{ TEX} + \beta_3 \text{ FEMALE} + \beta_4 \text{ AGE} \\ + \beta_5 \text{ NY} \cdot \text{AGE} + \beta_6 \text{ TEX} \cdot \text{AGE}$$

State	Age effect	Estimated age effect
California	β_4	0.005
New York	$\beta_4 + \beta_5$	0.005+0.010
Texas	$\beta_4 + \beta_6$	0.005-0.012

Test of $H_0 : \beta_5 = \beta_6 = 0$ gives Wald= 15.3 on 2 df, so strong evidence of interaction effect

$\hat{\pi}$ by Age and State



Note: Female fixed at 0.54 (sample proportion)