Why is Mixture Modelling so popular?

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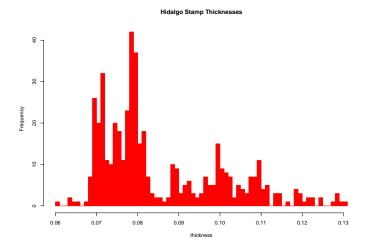
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- 2. Basic mixture formulation.
- 3. Ways to fit.
- 4. Inferential difficulties.
- 5. Cautionary example.

Introduction - Model Heterogeneity.

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Loosely, model *M* is a weighted mixture of component models $\{M_i\}$, if

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 w_i represents the proportion of the data "explained" by M_i .

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 w_i represents the proportion of the data "explained" by M_i . Example : Density estimation

$$f(\mathbf{x}_j) = \sum_{i=1}^g w_i f_i(\mathbf{x}_j),$$

 $f_i(\mathbf{x}_j)$ are "standard" densities and $0 \le w_i \le 1$, and $\sum_{i=1}^{g} w_i = 1$.

We could model the thicknesses as a mixture of Gaussian distributions.

$$f(\mathbf{x}_j) = \sum_{i=1}^{g} w_i \phi_i(\mathbf{x}_j, \boldsymbol{\mu}_i, \boldsymbol{\sigma}_i^2),$$

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But this raises an immediate question?

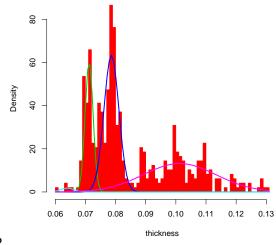
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But this raises an immediate question? Do we *know* how many component models there are in the mixture? Or, do we need to find out?

Four?

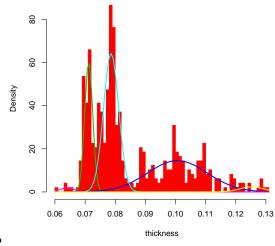
Stamp Thickness, g=4



Four?

Five?

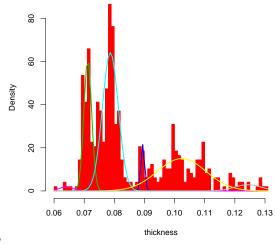
Stamp Thickness, g=5



Five?

Six?

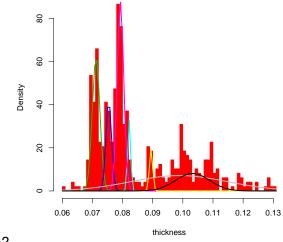
Stamp Thickness, g=6



Six?

Seven?

Stamp Thickness, g=7



Seven?

Model-Based Clustering

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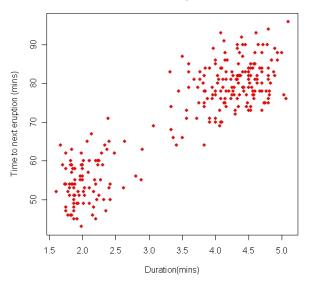
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Assumes clusters are individually well fitted by the models $\{M_i\}$.

Old Faithful Data



Old Faithful Eruption Data

Mixtures of regressions, linear Models, GLMs, survival,... EG. Hurn *et al.* (2000)

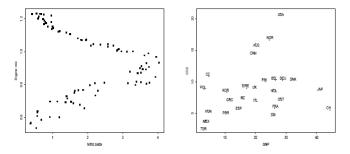
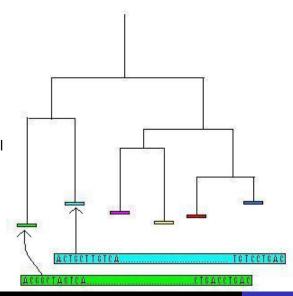


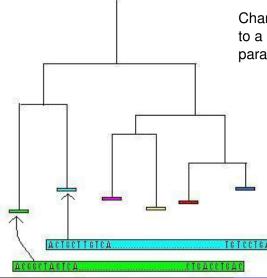
Fig. 2. (left) Equivalence ratio against exhaust nitric oxide concentration (Source: Hurvich et al., 1998); (right) representation of the GNP and CO2 emission levels in 1996 for various countries (Source: OECD).

Mixtures of Phylogenetic Models (Evolutionary tree of species)



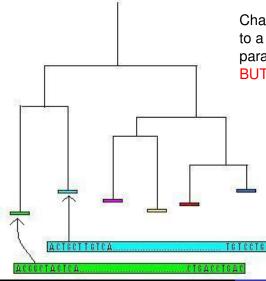
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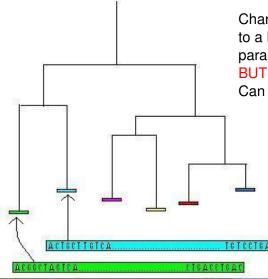
Characters (ACGT) evolve according to a Markov process Θ involving parameters such as mutation rates.

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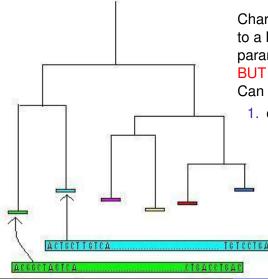
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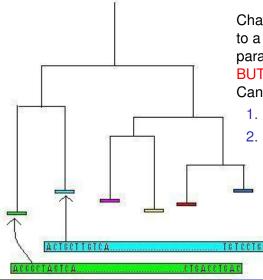
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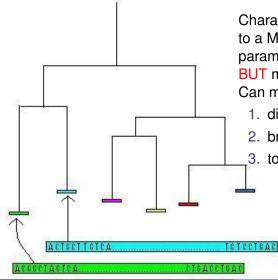
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- 3. topologies

Social Networks Handcock et al(2006)

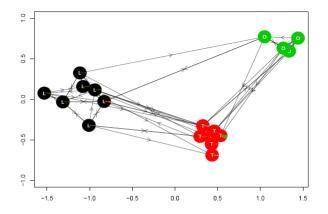


Figure 3: Estimates of clusters and latent positions for the relationship between monks within a monastery from the Bayesian estimation of the LPCM. The probability of assignment to each latent clusters is shown by a colored pie chart.

Two popular methods of fitting.

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1. EM Algorithm

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The EM algorithm involves assessing to which component j of the mixture each datum is expected to belong. Once this is established for all the data, fitting by usual MLE proceeds for the parameters of model j.

McMC may also utilise this device for completing the "missing" part of the data.

Allocation Variables

 Latent allocation variables describe which observations are assigned to each of the current components at each iteration of the EM algorithm or McMC sampler. In the case of McMC these provide a sample cluster configuration per McMC iteration.

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- Set $Z_j = i$ if observation *j* is allocated to component *i*
- Resulting allocation vector $\mathbf{Z} = (Z_1, ..., Z_n)$

Model selection, inference and interpretation can be complicated by several factors.

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- ▶ What is g?
- $g \text{ large} \Rightarrow \text{many parameters.}$
- Identifiability?

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- The likelihood

$$Lik = \prod_{i=1}^{n} \{ w_1 f(\mathbf{x}_i; \mathbf{\theta}_1) + \ldots + w_g f(\mathbf{x}_i; \mathbf{\theta}_g) \}$$

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- The term *label switching* is used to describe the invariance of the likelihood under the relabelling of the components.
- Often handled by imposing constraints, e.g label components in increasing order of weight w_j. Not always satisfactory.

Identifiability v Nonidentifiability

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► For example, consider the allocation vectors (4,4,3,3,4,2,2,3,1,3) and (2,2,1,1,2,3,3,1,4,1) these are different models if each component can be identified from some information.

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- ► However these two vectors define the same partition of the data and so are identical models from the clustering viewpoint. Therefore we would like some unique representation of them that identifies their common partition, i.e. {{1,2,5}, {3,4,8,10}, {6,7}, {9}}.

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Given allocation vectors from a sampler possibly suffering from label switching, we convert them to an unambiguous sample from space of possible partitions.

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- So estimating g as well as Gaussian parameters and weights.
- Standard approach to clustering is typically:
 - 1. Estimate likely number of components.
 - 2. Given *g*, estimate likely clustering.
 - 3. This conditional approach can easily mislead.

Old Faithful "Standard" Clustering

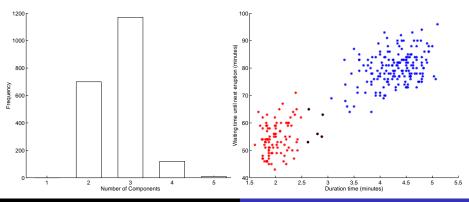
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The figure below shows the posterior distribution of the number of components k which has a prominent mode at 3. Alongside is the classification into 3 clusters obtained by standard hierarchical clustering using the number of times each pair of observations appear in the same sample cluster as a distance measure. (O'Hagan)

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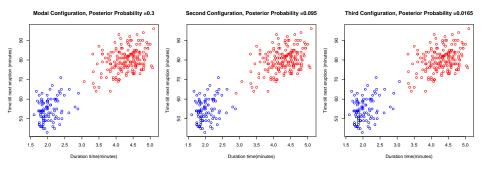
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Can we do more?

- Simple to explore the commonly occuring partitions.
- But what about variation in the sampled partition values?
- These live in a very complex and high dimensional discrete space.
- Could we get some insight into nature of this sample distribution?
- Dissimilarity between each pair of partitions number of pairs of observations that agree in the two partions, ie. either both in same group or both in different groups.

MDS plot of Old Faithful partitions

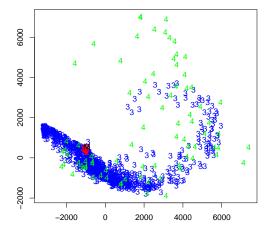
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isoMDS of unique sampled allocations

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- THANK YOU.